

# A simple effect size estimator for single case designs using WinBUGS

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March 13, 2012

Thanks to  
Institute of Education Sciences  
for support through grants:  
R305D100046  
R305D100033

# Usual Effect Size for Two Group Study

- ▶ Most common effect size (ES) measures compare means
- ▶ Because different studies use different scales, we must adjust for variance
- ▶ Thus we have the basic form  $ES = \frac{\bar{X}_T - \bar{X}_C}{s}$
- ▶ Notation: T = Treatment group, C = control group, s is standard deviation (of control or pooled)
- ▶ Several adjustments have been proposed to reduce bias

# Extension to Single Case Design (SCD)

- ▶ Why should we? After all, the dependent variable is often already on a scale that is directly interpretable.
- ▶ Because: Sometimes you want to compare SCDs with between-group studies that use different outcomes
- ▶ One approach: divide (treatment - baseline) by respondent's baseline standard deviation
- ▶ Why not? This is a within-person standard deviation
- ▶ Between-person variation (from two-group studies) is usually much larger, thus not comparable

# An Effect Size (ES) Comparable to Usual $d$

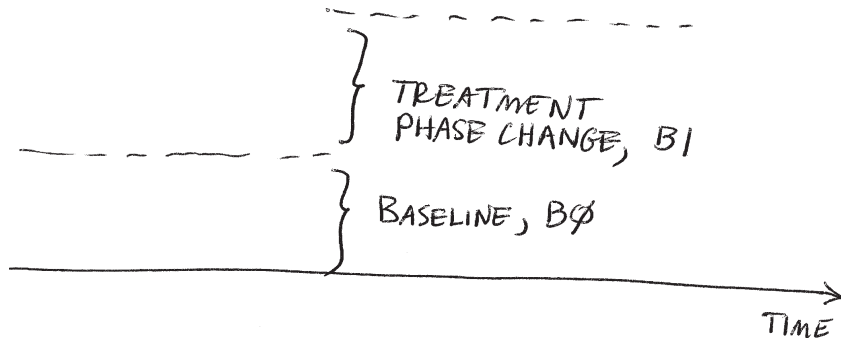
- ▶ Suppose an SCD has several participants
- ▶ Use variation among them as denominator
- ▶ Several options are possible, each with advantages and disadvantages
- ▶ Method 1: Divide by estimate of between-person variance (flawed, though it seems natural)
- ▶ Method 2: Divide by between + within variance (right, though not intuitive)

# Complication: Dependent variables are often counts, not continuous

- ▶ Ex 1: How many times does S hit another child?
  - ▶ No (theoretical) max
  - ▶ Poisson (or more complicated) distribution
  - ▶ Stat notation:  $Y \sim \text{Poisson}(\lambda)$
  - ▶ BUGS (Bayesian software) notation:  
`y[i] ~ dpois(lambda[i])`
- ▶ Ex 2: How many HW probs (out of 10) does S attempt?
  - ▶ attempt/10 = proportion
  - ▶ binomial (or more complicated) distribution
  - ▶ Stat notation:  $Y \sim \text{Binomial}(p, n)$
  - ▶ BUGS notation: `y[i] ~ dbin(p,n)`

# Graphical representation of important concepts

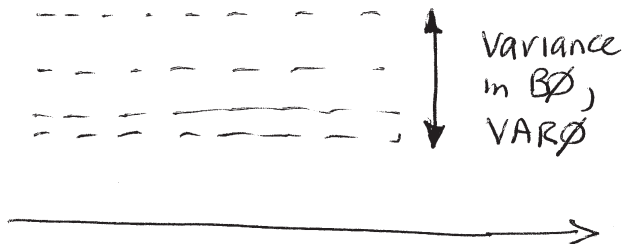
(1) Sketch of simple phase-change model for a single subject



Each subject has an average baseline value, and an average change during the treatment phase.

# Grahical representation

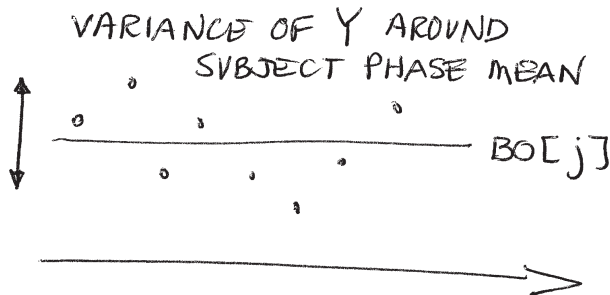
(2)Subjects vary in their baseline levels





## Graphical representation

(3) Within each subject, there is variation around the phase mean



# Statistical model for a continuous outcome and one person-level predictor

Model for individual observations for subject  $j$  at time  $i$ :

$$Y_{ij} = \beta_{0j} + \beta_{1j}Phase_{ij} + r_{ij}$$

$$Phase_{ij} = \begin{cases} 0 & \text{for baseline} \\ 1 & \text{for treatment} \end{cases}$$

We can model the baseline value of subject  $j$  as a function of Sex:

$$\beta_{0j} = \gamma_{00} + \gamma_{01}Female_j + u_{0j}$$

(where  $Female_j$  is 0 for males, 1 for females)

And also the treatment effect of subject  $j$  as a function of Sex:

$$\beta_{1j} = \gamma_{10} + \gamma_{11}Female_j + u_{1j}$$

## Simplified and combined statistical model for a continuous outcome (no person-level predictor)

Model for individual observations for subject  $j$  at time  $i$ :

$$Y_{ij} = \beta_{0j} + \beta_{1j}Phase_{ij} + r_{ij}$$

Model for baseline value of subject  $j$ :

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

Model for treatment effect of subject  $j$ :

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

Combined model, substituting into first equation:

$$Y_{ij} = (\gamma_{00} + u_{0j}) + (\gamma_{10} + u_{1j})Phase_{ij} + r_{ij}$$

We often use the combined model for certain software packages, including the BUGS software I will be using.

## More details on model

- ▶  $\text{Var}(u_{0j}) = \tau_{00}$  is part of the between-person variation
- ▶  $\text{Var}(r_{ij}) = \sigma^2$  is within person variation
  - ▶ contributes to usual denominator in ES
  - ▶ can't be separated from between-person variation without repeated measurements of the same person
- ▶ Thus, we might naively use  $ES_1 = \gamma_{10}/\sqrt{\tau_{00}}$
- ▶ But more properly we would define  $ES_2 = \gamma_{10}/\sqrt{\tau_{00} + \sigma^2}$
- ▶ On logit scale,  $\sigma^2 \approx 1/(n\pi(1 - \pi))$

# Logistic version of model

- ▶ Our data has number of successes in 10 trials, measured each day (or session) for the period of the study
- ▶ We need the equivalent of a logistic regression
- ▶ Outcome is the logarithm of the odds, called the logit
- ▶ In GLM (generalized linear models) this is expressed by separating the linear part of the model from the (logit) transform of the dependent variable

## Logistic model, cont.

We represent the combined model as follows:

$$\eta_{ij} = (\gamma_{00} + u_{0j}) + (\gamma_{10} + u_{1j})Phase_{ij}$$

$$\eta_{ij} = \ln\left(\frac{\pi_{ij}}{1-\pi_{ij}}\right) = \text{logit}(\pi)$$

$$Y_{ij} \sim \text{binomial}(\pi_{ij}, n_{ij})$$

Where  $Y_{ij}$  is the count of events out of  $n_{ij}$  trials,  
each with probability  $\pi_{ij}$  of the event occurring  
(Often  $n_{ij} = n$ ; that is, it is constant across time and subjects)

# Why Bayesian?

- ▶ Bayesian philosophy: Statistics is about using data to revise beliefs about unknown values (parameters)
- ▶ Initial (prior) beliefs could be vague (noninformative) or based on evidence available before collecting current data
- ▶ Information in the data is combined with any prior information to produce a Posterior Distribution, which summarizes our beliefs after seeing the data
- ▶ Bayesian models resemble usual models, except for specification of priors and interpretation of outcome
- ▶ Bayesian interpretation of interval: Probability is .95 that the parameter is in the interval (natural, but wrong in classical stat)

# Why (Win)BUGS?

- ▶ Some Bayesian computational methods, including those used in BUGS, allow simple ways to make inferences about derived quantities
- ▶ In particular, we want to make inferences about effect sizes, which are complicated quantities
- ▶ BUGS will make it (relatively) simple for us to make inferences about effect sizes
- ▶ The following slides will show how to write a simple logistic regression model, compute ES, and interpret results



# Data Structure

subj[ ]	r[ ]	phase[ ]			
			2	10	1
1	4	0	2	9	1
1	5	0	3	6	0
1	5	0	3	6	0
1	4	0	...		
1	5	0	3	9	1
1	10	1	3	9	1
1	8	1	4	4	0
....			4	3	0
1	9	1	...		
1	9	1	4	9	1
2	2	0	4	9	1
2	3	0	END	DATA	
...					

## WinBUGS Code 1: Basic Model

First some comments so I remember what I'm doing:

```
# binomial, 10 trials per session  
# 4 respondents, 2 phases (AB), multiple baseline  
# p(yes) goes from about .5 in baseline to .8 or .9
```

Next the model, first expressing the logit (log-odds) as a function of phase, then the distribution as a binomial:

```
model  
{ for (i in 1:103)  
  logit(p[i]) <- base[subj[i]] +  
    trt[subj[i]] * phase[i]  
  r[i] ~ dbin(p[i],10) }
```

Each subject has his/her own baseline and treatment effect, with mean  $\mu$  and precision ( $1/\text{variance}$ )  $\text{prec}$ :

```
for (j in 1:4)  
{ base[j] ~ dnorm(mu0, prec0)  
  trt[j] ~ dnorm(mu1, prec1) }
```

## WinBUGS Code 2: Priors and Create Variances, SDs

First, relatively uninformative priors for means and standard deviations of baseline and treatment effect:

```
mu0 ~ dnorm(0, .001)
mu1 ~ dnorm(0, .001)
prec0 ~ dgamma(.01, .01)
prec1 ~ dgamma(.01, .01)
```

Next, define variances and standard deviations, because it's hard for most of us to think in terms of precisions:

```
var0 <- 1/prec0
var1 <- 1/prec1
sd0 <- sqrt(var0)
sd1 <- sqrt(var1)
```

## WinBUGS Code 3: Create New Variables

Find estimate of within-person variation at baseline;  
must first transform from logit scale to find mean probability:

```
odds0 <- exp(mu0)
prob0 <- odds0/(1+odds0)
sigma.2 <- 1/(10 * prob0 * (1-prob0))
```

Next create total variance (between + within), and denominator for effect size estimate:

```
var.tot <- sigma.2 + var0
sd.tot  <- sqrt(var.tot)
```

Compute ES estimates, first wrong and then correctly:

```
# Uses only variation in average baselines:
es.bet.1  <- mu1/sd0

# (Properly) uses total variation:
es.bet.2  <- mu1/sd.tot
```

```
}
```

## Output: Effect Sizes 1

node	mean	sd	2.5%	median	97.5%
es.bet.1	5.5219	3.52	1.3147	4.8045	14.353
es.bet.2	2.7805	0.74798	1.195	2.8365	4.119

- ▶ First has a denominator that is too small, and therefore the ES estimate is too large
- ▶ Second should be (at least approximately) right

Advantages of MCMC(BUGS): Not just estimate, but also

- ▶ Standard error (called sd in output)
- ▶ CI (2.5% – 97.5%)
- ▶ Info on skewness of distribution

## Output: Effect Sizes 2

Rounded estimates for ES2 (correct estimate):

node	mean	sd	2.5%	median	97.5%
<hr/>					
es.bet.2	2.78	0.75	1.20	2.84	4.12

For ES2, we have

- ▶ Evidence of minor skewness:  
Compare  $2.84 - 1.20 = 1.64$  to  $4.12 - 2.84 = 1.28$
- ▶ More evidence skew is minor:  
Mean (2.78) is close to median (2.84)
- ▶ Empirical 95 percent credible (confidence) interval:  
(1.20, 4.12)
- ▶ Thus ES could be as small as about 1, or as large as about 4
- ▶ Quite wide interval due to small number of respondents

## Output : Basic Parameters

node	mean	sd	2.5%	median	97.5%
mu0	-0.099319	0.38842	-0.84193	-0.10322	0.66174
mu1	2.3425	0.36558	1.6788	2.3369	3.0548
sd0	0.59467	0.44389	0.17013	0.48249	1.7053
sd1	0.47212	0.43196	0.089285	0.36225	1.5825

- ▶ Baseline average odds :  $\exp(-.099) = .905$
- ▶ Baseline average proportion:  $.905/1.905 = .475$
- ▶ Treatment phase average odds :  $\exp(-.909 + 2.343) = 9.42$
- ▶ Treatment phase average proportion :  $9.42/10.42 = .904$

## Output : Variances

Variances among subjects in baseline log-odds  
and treatment effects:

node	mean	sd	2.5%	median	97.5%
var0	0.55067	1.7968	0.028945	0.2328	2.9079
var1	0.40949	1.4639	0.007972	0.13123	2.5044

Variance within phases for subjects (sigma.2) and total (var.tot):

sigma.2	0.42011	0.15086	0.40001	0.40405	0.50187
var.tot	0.97079	1.9037	0.43176	0.64223	3.3746
sd.tot	0.91024	0.37715	0.65708	0.80139	1.837



# Estimate for Individuals

Baseline log-odds for each person:

node	mean	sd	2.5%	median	97.5%
-----					
base[1]	0.032035	0.18773	-0.33265	0.032001	0.40112
base[2]	-0.39729	0.20606	-0.81881	-0.39097	-0.01188
base[3]	0.40382	0.21969	-0.02369	0.40348	0.83808
base[4]	-0.44905	0.20319	-0.84266	-0.44938	-0.04398

Treatment effect (on log-odds scale) for each person:

trt[1]	2.5287	0.30277	1.9984	2.5041	3.1901
trt[2]	2.5158	0.2903	1.9742	2.5042	3.1398
trt[3]	2.3207	0.30109	1.7497	2.3133	2.9469
trt[4]	1.9931	0.27734	1.4186	2.006	2.4951

# Problems with this approach

- ▶ Standardized measures aren't in original scale
- ▶ With small number of observations in any phase, the difference between phases is not well-estimated
- ▶ With small number of respondents, standard deviation among respondents is not well-estimated
- ▶ If respondents are selected for low (or high) initial status, between-respondent variation may be artificially low compared to between-group studies (needs checking, at least)

# Conclusions

- ▶ Bayesian computation allows relatively simple
  - ▶ Estimation of effect size
  - ▶ Production of confidence interval (CI)
  - ▶ Estimates for each individual, including CI
- ▶ Some additional training (beyond HLM) is needed to set up the model, do computations, and interpret results
- ▶ This procedure should be useful for SCD researchers who want to produce results comparable to between-group studies

# References

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